

# Rosenbaum-Style p-Values for Matched Observational Studies with Unmeasured Confounding

Hanyu Wu

Department of Biostatistics, Peking University

2024.3.14



- ① Matched Data Sensitivity Model
- ② FRT under Rosenbaum's Sensitivity Model
- ③ Application

- 1 Matched Data Sensitivity Model
- 2 FRT under Rosenbaum's Sensitivity Model
- 3 Application

# Model Setting

- Consider **exactly matched pairs** from an observational study, with  $(i, j)$  indexing unit  $j$  in pair  $i$  ( $i = 1, \dots, n; j = 1, 2$ ).
- In pair  $i$ , we have  $X_{i1} = X_{i2} = X_i$ . ( $e(X_{i1}) = e(X_{i2})$ )

# Model Setting

- Consider **exactly matched pairs** from an observational study, with  $(i, j)$  indexing unit  $j$  in pair  $i$  ( $i = 1, \dots, n; j = 1, 2$ ).
- In pair  $i$ , we have  $X_{i1} = X_{i2} = X_i$ . ( $e(X_{i1}) = e(X_{i2})$ )
- Assume iid sampling, and define the propensity score as

$$e_{ij} = \text{pr} \{Z_{ij} = 1 \mid X_i, Y_{ij}(1), Y_{ij}(0)\}.$$

# Model Setting

- Consider **exactly matched pairs** from an observational study, with  $(i, j)$  indexing unit  $j$  in pair  $i$  ( $i = 1, \dots, n; j = 1, 2$ ).
- In pair  $i$ , we have  $X_{i1} = X_{i2} = X_i$ . ( $e(X_{i1}) = e(X_{i2})$ )
- Assume iid sampling, and define the propensity score as

$$e_{ij} = \text{pr} \{Z_{ij} = 1 \mid X_i, Y_{ij}(1), Y_{ij}(0)\}.$$

- Let  $\mathbb{S}_i = \{Y_{i1}(1), Y_{i1}(0), Y_{i2}(1), Y_{i2}(0)\}$  denote the set of all potential outcomes within pair  $i$ .

# Conditional Assignment Probability

Conditioning on the event that  $Z_{i1} + Z_{i2} = 1$ , consider

$$\pi_{i1} = \text{pr} \{ Z_{i1} = 1 \mid X_i, \mathbb{S}_i, Z_{i1} + Z_{i2} = 1 \}.$$

# Conditional Assignment Probability

Conditioning on the event that  $Z_{i1} + Z_{i2} = 1$ , consider

$$\pi_{i1} = \text{pr} \{Z_{i1} = 1 \mid X_i, S_i, Z_{i1} + Z_{i2} = 1\}.$$

Define  $o_{ij} = e_{ij} / (1 - e_{ij})$  as the odds of the treatment for unit  $(i, j)$ , we have

$$\begin{aligned} \pi_{i1} &= \frac{\text{pr} \{Z_{i1} = 1, Z_{i2} = 0 \mid X_i, S_i\}}{\text{pr} \{Z_{i1} + Z_{i2} = 1 \mid X_i, S_i\}} \\ &= \frac{\text{pr} \{Z_{i1} = 1, Z_{i2} = 0 \mid X_i, S_i\}}{\text{pr} \{Z_{i1} = 1, Z_{i2} = 0 \mid X_i, S_i\} + \text{pr} \{Z_{i1} = 0, Z_{i2} = 1 \mid X_i, S_i\}} \\ &= \frac{e_{i1} (1 - e_{i2})}{e_{i1} (1 - e_{i2}) + (1 - e_{i1}) e_{i2}} = \frac{o_{i1}}{o_{i1} + o_{i2}}. \end{aligned}$$



# With or Without Ignorability

- With ignorability holds,  $e_{ij}$  is only a function of  $X_i$ , and therefore,  $e_{i1} = e_{i2}$  and  $\pi_{i1} = 1/2$ .

## With or Without Ignorability

- With ignorability holds,  $e_{ij}$  is only a function of  $X_i$ , and therefore,  $e_{i1} = e_{i2}$  and  $\pi_{i1} = 1/2$ .
- The treatment assignment is identical to the MPE conditioning on the covariates and the event that each pair has a treated and control units.

# With or Without Ignorability

- With ignorability holds,  $e_{ij}$  is only a function of  $X_i$ , and therefore,  $e_{i1} = e_{i2}$  and  $\pi_{i1} = 1/2$ .
- The treatment assignment is identical to the MPE conditioning on the covariates and the event that each pair has a treated and control units.
- So we can analyze the exactly matched observational study as if it is a MPE, **using either the FRT or the Neymanian approach in Chapter 7.**

## With or Without Ignorability

- Without ignorability,  $e_{ij} \not\perp \{Y_{ij}(1), Y_{ij}(0)\} | X_i$ , so  $o_{i1} \neq o_{i2}$ . We introduce Rosenbaum (1987b)'s model for sensitivity analysis imposes bounds on the odds ratio  $o_{i1}/o_{i2}$ .

# With or Without Ignorability

- Without ignorability,  $e_{ij} \not\perp \{Y_{ij}(1), Y_{ij}(0)\} | X_i$ , so  $o_{i1} \neq o_{i2}$ . We introduce Rosenbaum (1987b)'s model for sensitivity analysis imposes bounds on the odds ratio  $o_{i1}/o_{i2}$ .

## Assumption 19.1 (Rosenbaum's sensitivity model)

The odds ratios are bounded by

$$o_{i1}/o_{i2} \leq \Gamma, \quad o_{i2}/o_{i1} \leq \Gamma$$

for some pre-specified  $\Gamma \geq 1$ . Equivalently,

$$\frac{1}{1+\Gamma} \leq \pi_{i1} \leq \frac{\Gamma}{1+\Gamma}$$

for some pre-specified  $\Gamma \geq 1$ .

# Rosenbaum's sensitivity model

## Assumption 19.1 (Rosenbaum's sensitivity model)

The odds ratios are bounded by

$$o_{i1}/o_{i2} \leq \Gamma, \quad o_{i2}/o_{i1} \leq \Gamma$$

for some pre-specified  $\Gamma \geq 1$ . Equivalently,

$$\frac{1}{1+\Gamma} \leq \pi_{i1} \leq \frac{\Gamma}{1+\Gamma}$$

for some pre-specified  $\Gamma \geq 1$ .

- Under Assumption 19.1, we have a biased MPE with unequal and varying treatment and control probabilities across pairs.

# Rosenbaum's sensitivity model

## Assumption 19.1 (Rosenbaum's sensitivity model)

The odds ratios are bounded by

$$o_{i1}/o_{i2} \leq \Gamma, \quad o_{i2}/o_{i1} \leq \Gamma$$

for some pre-specified  $\Gamma \geq 1$ . Equivalently,

$$\frac{1}{1+\Gamma} \leq \pi_{i1} \leq \frac{\Gamma}{1+\Gamma}$$

for some pre-specified  $\Gamma \geq 1$ .

- Under Assumption 19.1, we have a biased MPE with unequal and varying treatment and control probabilities across pairs.
- When  $\Gamma = 1$ , we have  $\pi_{i1} = \frac{1}{2}$  and thus a standard MPE.

# Rosenbaum's sensitivity model

## Assumption 19.1 (Rosenbaum's sensitivity model)

The odds ratios are bounded by

$$o_{i1}/o_{i2} \leq \Gamma, \quad o_{i2}/o_{i1} \leq \Gamma$$

for some pre-specified  $\Gamma \geq 1$ . Equivalently,

$$\frac{1}{1+\Gamma} \leq \pi_{i1} \leq \frac{\Gamma}{1+\Gamma}$$

for some pre-specified  $\Gamma \geq 1$ .

- Under Assumption 19.1, we have a biased MPE with unequal and varying treatment and control probabilities across pairs.
- When  $\Gamma = 1$ , we have  $\pi_{i1} = \frac{1}{2}$  and thus a standard MPE.
- $\Gamma > 1$  measures the deviation from the ideal MPE due to the **omitted variables in matching**.



- 1 Matched Data Sensitivity Model
- 2 FRT under Rosenbaum's Sensitivity Model
- 3 Application

# FRT Setting

- Consider testing the sharp null hypothesis

$$H_{0F} : Y_{ij}(1) = Y_{ij}(0) \text{ for } i = 1, \dots, n \text{ and } j = 1, 2$$

based on within-pair differences  $\hat{\tau}_i = (2Z_{i1} - 1)(Y_{i1} - Y_{i2})$   
( $i = 1, \dots, n$ ).

# FRT Setting

- Consider testing the sharp null hypothesis

$$H_{0F} : Y_{ij}(1) = Y_{ij}(0) \text{ for } i = 1, \dots, n \text{ and } j = 1, 2$$

based on within-pair differences  $\hat{\tau}_i = (2Z_{i1} - 1)(Y_{i1} - Y_{i2})$   
( $i = 1, \dots, n$ ).

- Under  $H_{0F}$ ,  $|\hat{\tau}_i|$  is fixed but  $S_i = I(\hat{\tau}_i > 0)$  is random if  $\hat{\tau}_i \neq 0$ .

# FRT Setting

- Consider testing the sharp null hypothesis

$$H_{0F} : Y_{ij}(1) = Y_{ij}(0) \text{ for } i = 1, \dots, n \text{ and } j = 1, 2$$

based on within-pair differences  $\hat{\tau}_i = (2Z_{i1} - 1)(Y_{i1} - Y_{i2})$   
( $i = 1, \dots, n$ ).

- Under  $H_{0F}$ ,  $|\hat{\tau}_i|$  is fixed but  $S_i = I(\hat{\tau}_i > 0)$  is random if  $\hat{\tau}_i \neq 0$ .
  - Specifically,  $\hat{\tau}_i$  takes value from  $Y_{i1}(1) - Y_{i2}(0)$  and  $Y_{i2}(1) - Y_{i1}(0)$ . Under  $H_{0F}$ ,  $|\hat{\tau}_i|$  is constant.

# Test Statistics

- Consider the following class of test statistics:

$$T = \sum_{i=1}^n S_i q_i$$

where  $q_i \geq 0$  is a function of  $(|\hat{\tau}_1|, \dots, |\hat{\tau}_n|)$ .

# Test Statistics

- Consider the following class of test statistics:

$$T = \sum_{i=1}^n S_i q_i$$

where  $q_i \geq 0$  is a function of  $(|\hat{\tau}_1|, \dots, |\hat{\tau}_n|)$ .

- Special cases include the sign statistic, the pair  $t$  statistic (up to some constant shift), and the Wilcoxon sign rank statistic:

$$T_1 = \sum_{i=1}^n S_i, \quad T_2 = \sum_{i=1}^n S_i |\hat{\tau}_i|, \quad T_3 = \sum_{i=1}^n S_i R_i,$$

where  $(R_1, \dots, R_n)$  are the ranks of  $(|\hat{\tau}_1|, \dots, |\hat{\tau}_n|)$ .

# Null Distribution of Test Statistic

- $E(S_i) = \pi_{i1} \neq \frac{1}{2}$  without ignorability assumption, so the null distribution of  $T$  is unspecified.

# Null Distribution of Test Statistic

- $E(S_i) = \pi_{i1} \neq \frac{1}{2}$  without ignorability assumption, so the null distribution of  $T$  is unspecified.
- Among all null distributions, we aim to find "the worst case", under which  $T$  has the largest  $p$ -value.



# Null Distribution of Test Statistic

- $E(S_i) = \pi_{i1} \neq \frac{1}{2}$  without ignorability assumption, so the null distribution of  $T$  is unspecified.
- Among all null distributions, we aim to find "the worst case", under which  $T$  has the largest  $p$ -value.
- Since FRT is right-tailed test, the worst case is  $E(S_i) = \max(\pi_{i1})$ , i.e.

$$S_i \stackrel{\text{IID}}{\sim} \text{Bernoulli} \left( \frac{\Gamma}{1 + \Gamma} \right).$$

# Null Distribution of Test Statistic

- Under "the worst case",

$$E_{\Gamma}(T) = \frac{\Gamma}{1 + \Gamma} \sum_{i=1}^n q_i, \quad \text{var}_{\Gamma}(T) = \frac{\Gamma}{(1 + \Gamma)^2} \sum_{i=1}^n q_i^2$$

# Null Distribution of Test Statistic

- Under "the worst case",

$$E_{\Gamma}(T) = \frac{\Gamma}{1+\Gamma} \sum_{i=1}^n q_i, \text{ var}_{\Gamma}(T) = \frac{\Gamma}{(1+\Gamma)^2} \sum_{i=1}^n q_i^2$$

- The Normal approximation is

$$\frac{T - \frac{\Gamma}{1+\Gamma} \sum_{i=1}^n q_i}{\sqrt{\frac{\Gamma}{(1+\Gamma)^2} \sum_{i=1}^n q_i^2}} \xrightarrow{d} N(0, 1).$$

# Null Distribution of Test Statistic

- Under "the worst case",

$$E_{\Gamma}(T) = \frac{\Gamma}{1+\Gamma} \sum_{i=1}^n q_i, \quad \text{var}_{\Gamma}(T) = \frac{\Gamma}{(1+\Gamma)^2} \sum_{i=1}^n q_i^2$$

- The Normal approximation is

$$\frac{T - \frac{\Gamma}{1+\Gamma} \sum_{i=1}^n q_i}{\sqrt{\frac{\Gamma}{(1+\Gamma)^2} \sum_{i=1}^n q_i^2}} \xrightarrow{d} N(0, 1).$$

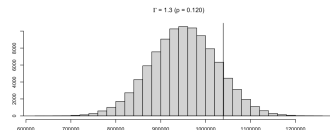
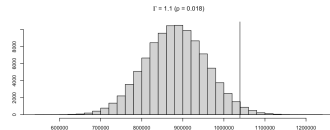
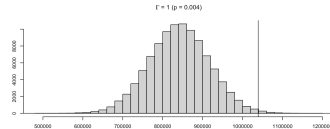
- In practice, we can report a sequence of  $p$ -values as a function of  $\Gamma$ .

- 1 Matched Data Sensitivity Model
- 2 FRT under Rosenbaum's Sensitivity Model
- 3 Application**

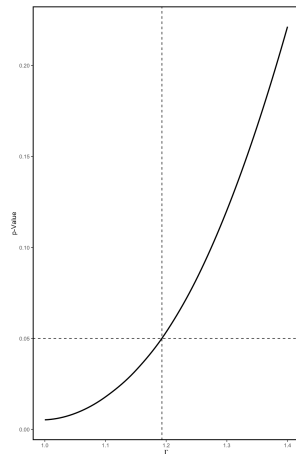
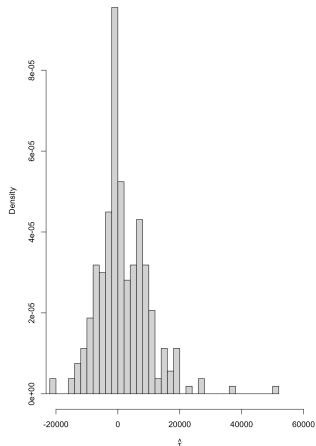
# Matched Lalonde Data Analysis

```
# Match
library("car")
library("Matching")
data(lalonde)
lalonde$u74 <- as.numeric(lalonde$re74 == 0)
lalonde$u75 <- as.numeric(lalonde$re75 == 0)
y = lalonde$re78
z = lalonde$treat
x = as.matrix(lalonde[, c("age", "educ", "black",
                          "hisp", "married", "nodegr",
                          "re74", "re75", "u74", "u75")])
matchest.adj = Match(Y = y, Tr = z, X = x, BiasAdjust = TRUE)
```

```
# Rosenbaum's Sensitivity Analysis
tre = lalonde$re78[matchest.adj$index.treated]
con = lalonde$re78[matchest.adj$index.control]
tau = tre - con
T = sum(tau*(tau>0))
ga = 1 # Gamma = 1
Mean1 = ga/(1+ga)*sum(abs(tau))
Var1 = ga/(1+ga)^2*sum((tau)^2)
hist(rnorm(100000, Mean1, sqrt(Var1)),
     main = expression(paste(Gamma, " = 1 (p = 0.004)")),
     xlab = NULL, ylab = NULL)
abline(h=0,v=T)p1 = pnorm(T, Mean1, sqrt(Var1))
```



# Matched Lalonde Data Analysis



*Thanks!*